

# MATH/STAT 355: Problem Set 3

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## Properties of Estimators

1. Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Exponential}(\lambda)$ . Find a sufficient statistic for  $\lambda$ .

2. In this problem, we'll prove that

$$\left( \sqrt{\frac{n-1}{2}} \right) \frac{\Gamma(\frac{n-1}{2})}{\Gamma(\frac{n}{2})} S$$

is the UMVUE for  $\sigma$ , where  $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$  and  $X_i \stackrel{iid}{\sim} N(\mu, \sigma^2)$ .

(a) Show that  $\sum_{i=1}^n (X_i - \bar{X})^2 = \sum_{i=1}^n (X_i - \mu)^2 - n(\bar{X} - \mu)^2$ .

(b) Show that  $(\bar{X}, S^2)$  is a jointly sufficient statistic for  $(\mu, \sigma^2)$ .

(c) Show that  $\frac{(n-1)S^2}{\sigma^2} = \sum_{i=1}^n \left( \frac{X_i - \mu}{\sigma} \right)^2 - \left( \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \right)^2$

(d) Show that  $\sum_{i=1}^n \left( \frac{X_i - \mu}{\sigma} \right)^2 \sim \chi_n^2$ , and  $\left( \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \right)^2 \sim \chi_1^2$ . Use the known relationships that if  $X \sim N(0, 1)$ , then  $X^2 \sim \chi_1^2$ , and that if, for  $Y \perp W$ ,  $Y \sim \chi_n^2$  and  $W \sim \chi_m^2$ ,  $Y + W \sim \chi_{n+m}^2$ .

(e) Using everything we've done in the previous steps, show (finally) that  $\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$ .

(f) Let  $Y = \frac{(n-1)S^2}{\sigma^2}$ , and note that  $\sqrt{\frac{\sigma^2 Y}{n-1}} = S$ . You may take without proof that  $(\bar{X}, S^2)$  is a *complete*, jointly sufficient statistic for  $(\mu, \sigma^2)$ . Use RBLs to show that  $\left( \sqrt{\frac{n-1}{2}} \right) \frac{\Gamma(\frac{n-1}{2})}{\Gamma(\frac{n}{2})} S$  is the UMVUE for  $\sigma$ .

## Consistency

1. Prove Chebyshev's Corollary's

(a) Corollary 1: If  $\hat{\theta}_n$  is an unbiased estimator for  $\theta$  and  $\lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_n) = 0$ , then  $\hat{\theta}_n$  is consistent for  $\theta$ .

(b) Markov's Inequality, similar to Chebyshev's Inequality, states that if  $X$  is a non-negative random variable (i.e.,  $X \geq 0$ ), then  $P(X \geq a) \leq \frac{E(X)}{a}$  for all  $a > 0$ . Use this inequality, along with the squeeze/sandwich theorem and the relationship between mean-squared error  $E[(\hat{\theta}_n - \theta)^2]$ , variance, and bias to prove Chebyshev's Corollary 2: If  $\hat{\theta}_n$  is an *asymptotically* unbiased estimator for  $\theta$  and  $\lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_n) = 0$ , then  $\hat{\theta}_n$  is consistent for  $\theta$ .

2. Let  $Y_1, \dots, Y_n$  be a random sample of size  $n$  from a Normal pdf where  $E(Y_i) = 0$  and variance  $\text{Var}(Y_i) = \sigma^2$ . Show that  $S_n^2 = \frac{1}{n} \sum_{i=1}^n Y_i^2$  is a consistent estimator for  $\sigma^2$ .

## Correlated Data: Part 1

Consider the following autoregressive model

$$X_n = \beta X_{n-1} + e_n, \quad n = 1, 2, 3, \dots$$

with  $-1 < \beta < 1$ ,  $X_0 = 0$ , and  $e_1, \dots, e_n$  i.i.d. with  $E[e_i] = 0$  and  $Var[e_i] = 1$  (Note that we have not assumed a pdf for  $e_i$ , so in this problem, this process is actually *semiparametric*). Let  $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$  and  $\bar{e}_n = \frac{1}{n} \sum_{i=1}^n e_i$ .

If you've taken Correlated Data, you've seen this before! It's like an AR1! Wow, connections!

1. Calculate  $Var(X_n)$ .
2. Use Chebyshev's Inequality to show that  $X_n/\sqrt{n} \xrightarrow{p} 0$ . (Hint:  $\sum_{i=0}^{n-1} \beta^{2i}$  is a convergent geometric series).

We'll do more with this problem on the next Problem Set!